Clustering / Unsupervised Methods

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Introduction

- Until now, we've assumed our training samples are "labeled" by their category membership.
- Methods that use labeled samples are said to be *supervised*; otherwise, they're said to be *unsupervised*.
- However:
 - Why would one even be interested in learning with unlabeled samples?
 - Is it even possible in principle to learn anything of value from unlabeled samples?

- Collecting and labeling a large set of sample patterns can be surprisingly costly.
 - E.g., videos are virtually free, but accurately *labeling* the video pixels is expensive and time consuming.

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 - Train a classifier on a small set of samples, then tune it up to make it run without supervision on a large, unlabeled set.
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- It detect the gradual change of pattern over time.
- To find features that will then be useful for categorization.
- To gain insight into the nature or structure of the data during the early stages of an investigation.

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Source: A. K. Jain and R. C. Dubes. Alg. for Clustering Data, Prentiice Hall, 1988.

• What is data clustering?

- Grouping of objects into meaningful categories
- Given a representation of N objects, find k clusters based on a measure of similarity.

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 - Natural Classification: degree of similarity among forms.
 - Data exploration: discover underlying structure, generate hypotheses, detect anomalies.
 - Compression: for organizing data.
 - Applications: can be used by any scientific field that collects data!

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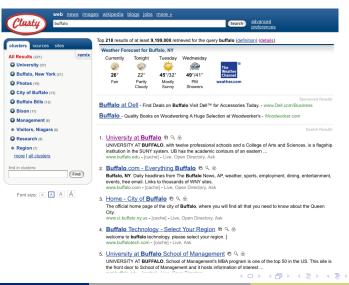
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 - Applications: can be used by any scientific field that collects data!
- Google Scholar: 1500 clustering papers in 2007 alone!

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E.g.: Structure Discovering via Clustering

Source: http://clusty.com



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E.g.: Topic Discovery Source: Map of Science, Nature, 2006

• 800,000 scientific papers clustered into 776 topics based on how often the papers were cited together by authors of other papers



Data Clustering - Formal Definition

• Given a set of N unlabeled examples $D = x_1, x_2, ..., x_N$ in a *d*-dimensional feature space, D is partitioned into a number of disjoint subsets D_j 's:

$$D = \cup_{j=1}^{k} D_j$$
 where $D_i \cap D_j = \emptyset, i \neq j$, (1)

where the points in each subset are similar to each other according to a given criterion ϕ .

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• A partition is denoted by

$$\pi = (D_1, D_2, ..., D_k)$$
(2)

and the problem of data clustering is thus formulated as

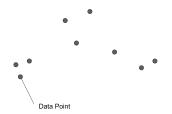
$$\pi^* = \underset{\pi}{\operatorname{argmin}} f(\pi) \quad , \tag{3}$$

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where $f(\cdot)$ is formulated according to ϕ .

Source: D. Aurthor and S. Vassilvitskii. *k*-Means++: The Advantages of Careful Seeding

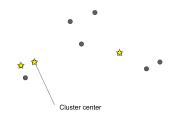
- Randomly initialize $\mu_1, \mu_2, ..., \mu_c$
- Repeat until no change in μ_i :
 - Classify N samples according to nearest μ_i
 - Recompute μ_i



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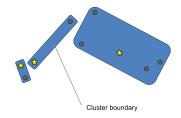


First choose k arbitrary centers

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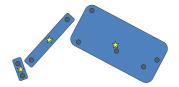


Assign points to closest centers

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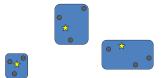


Recompute centers

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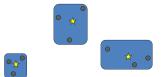


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Source: D. Aurthor and S. Vassilvitskii. *k*-Means++: The Advantages of Careful Seeding

- Choose starting centers iteratively.
- Let D(x) be the distance from x to the nearest existing center, take x as new center with probability $\propto D(x)^2$.
- Repeat until no change in μ_i :
 - Classify N samples according to nearest μ_i
 - Recompute μ_i
- (refer to the slides by D. Aurthor and S. Vassolvitskii for details)

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User's Dilemma Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- What is a cluster?
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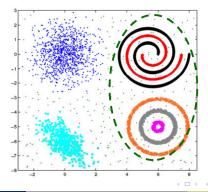
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- Are the discovered clusters and partition valid?

User's Dilemma

- What is a cluster?
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- Which features and normalization scheme?
- How many clusters?
- Which clustering method?
- O Are the discovered clusters and partition valid?
- O Does the data have any clustering tendency?

Cluster Similarity?

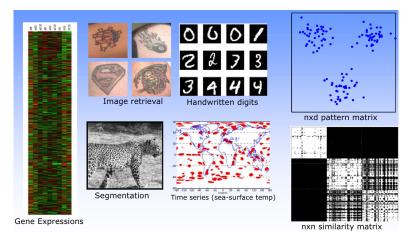
- Compact Clusters
 - Within-cluster distance < between-cluster connectivity
- Connected Clusters
 - Within-cluster connectivity > between-cluster connectivity
- Ideal cluster: compact and isolated.



Representation (features)?

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

• There's no universal representation; they're domain dependent.

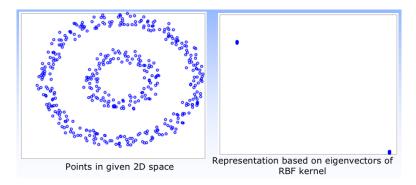


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Good Representation

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

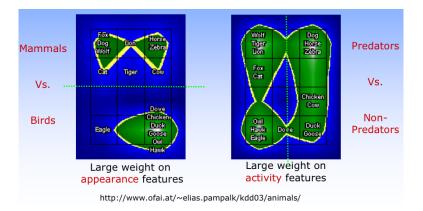
• A good representation leads to compact and isolated clusters.



How do we weigh the features?

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

• Two different meaningful groupings produced by different weighting schemes.

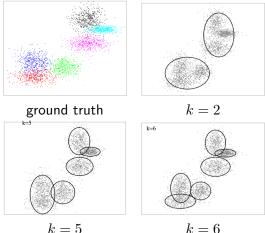


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How do we decide the Number of Clusters?

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

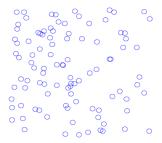
• The samples are generated by 6 independent classes, yet:

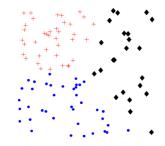


Cluster Validity

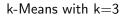
Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

• Clustering algorithms find clusters, even if there are no **natural** clusters in the data.





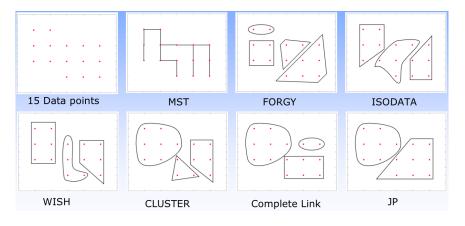
100 2D uniform data points



Comparing Clustering Methods

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

• Which clustering algorithm is the best?

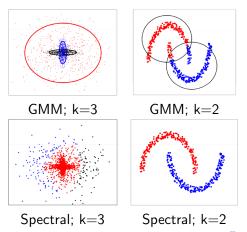


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There's no best Clustering Algorithm!

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- Each algorithm imposes a structure on data.
- Good fit between model and data \Rightarrow success.



• Recall the Gaussian distribution:

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{d/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}} \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right] \quad (4)$$

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- The Gaussian mixture is a **linear superposition of Gaussians** in the form:

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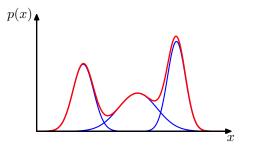
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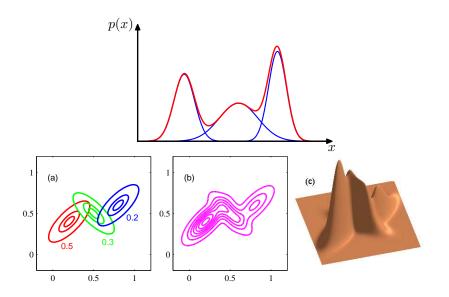
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 The π_k are non-negative scalars called mixing coefficients and they govern the relative importance between the various Gaussians in the mixture density. Σ_k π_k = 1.



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Introducing Latent Variables

• Define a *K*-dimensional binary random variable **z**.

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Introducing Latent Variables

- Define a *K*-dimensional binary random variable z.
- z has a 1-of-K representation such that a particular element z_k is 1 and all of the others are zero. Hence:

$$z_k \in \{0, 1\}$$

$$\sum_k z_k = 1$$
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• The marginal distribution over z is specified in terms of the mixing coefficients:

$$p(z_k = 1) = \pi_k \tag{8}$$

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And, recall, $0 \le \pi_k \le 1$ and $\sum_k \pi_k = 1$.

• Since ${\bf z}$ has a 1-of-K representation, we can also write this distribution as

$$p(\mathbf{z}) = \prod_{k=1}^{K} \pi_k^{z_k}$$

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$$p(\mathbf{z}) = \prod_{k=1}^{K} \pi_k^{z_k} \tag{9}$$

• The conditional distribution of \mathbf{x} given \mathbf{z} is a Gaussian:

$$p(\mathbf{x}|z_k = 1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$
(10)

or

$$p(\mathbf{x}|\mathbf{z}) = \prod_{k=1}^{K} \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_k}$$
(11)

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• We are interested in the marginal distribution of x:

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- So, given our latent variable z, the marginal distribution of x is a Gaussian mixture.
- If we have N observations $\mathbf{x}_1, \ldots, \mathbf{x}_N$, then because of our chosen representation, it follows that we have a latent variable \mathbf{z}_n for each observed data point \mathbf{x}_n .

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- γ(z_k) can also be viewed as the responsibility that component k takes
 for explaining the observation x.

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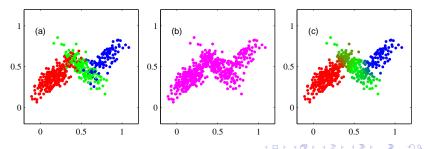
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• Ultimately, we want to find the values of the parameters π, μ, Σ that maximize this function.

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• However, maximizing the log-likelihood terms for GMMs is much more complicated than for the case of a single Gaussian. Why?

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- However, maximizing the log-likelihood terms for GMMs is much more complicated than for the case of a single Gaussian. Why?
- The difficulty arises from the sum over k inside of the log-term. The log function no longer acts directly on the Gaussian, and no closed-form solution is available.

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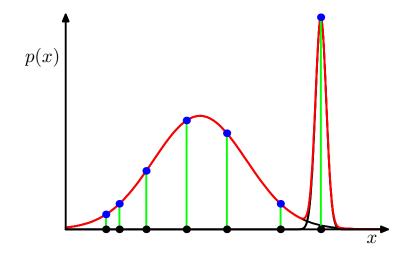
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- Consider the limit $\sigma_j \to 0$ to see that this term goes to infinity and hence the log-likelihood will also go to infinity.
- Thus, the maximization of the log-likelihood function is not a well posed problem because such a singularity will occur whenever one of the components collapses to a single, specific data point.



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• Note the natural appearance of the responsibility terms on the RHS.

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• Multiplying by $\mathbf{\Sigma}_k^{-1}$, which we assume is non-singular, gives

$$\boldsymbol{\mu}_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{x}_{n}$$
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- We see the $k^{\rm th}$ mean is the weighted mean over all of the points in the dataset.
- Interpret N_k as the number of points assigned to component k.
- We find a similar result for the covariance matrix:

$$\boldsymbol{\Sigma}_{k} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) (x_{n} - \boldsymbol{\mu}_{k}) (x_{n} - \boldsymbol{\mu}_{k})^{\mathsf{T}} \quad .$$
 (24)

• We also need to maximize $\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$ with respect to the mixing coefficients π_k .

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- We also need to maximize $\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$ with respect to the mixing coefficients π_k .
- Introduce a Lagrange multiplier to enforce the constraint $\sum_k \pi_k = 1$.

$$\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) + \lambda \left(\sum_{k=1}^{K} \pi_k - 1\right)$$
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• Eliminate λ and rearrange to obtain:

$$\pi_k = \frac{N_k}{N} \tag{28}$$

• So, we're done, right? We've computed the maximum likelihood solutions for each of the unknown parameters.

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- Wrong!

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- Wrong!
- The responsibility terms depend on these parameters in an intricate way:

$$\gamma(z_k) \doteq p(z_k = 1 | \mathbf{x}) = \frac{\pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

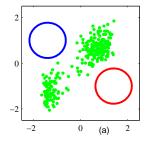
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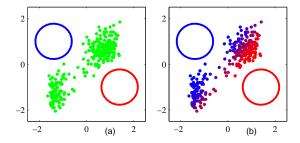
- But, these results do suggest an iterative scheme for finding a solution to the maximum likelihood problem.
 - **(**) Chooce some initial values for the parameters, π, μ, Σ .
 - Use the current parameters estimates to compute the posteriors on the latent terms, i.e., the responsibilities.
 - **③** Use the responsibilities to update the estimates of the parameters.
 - Repeat 2 and 3 until convergence.

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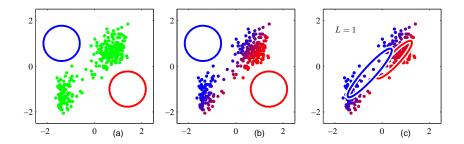
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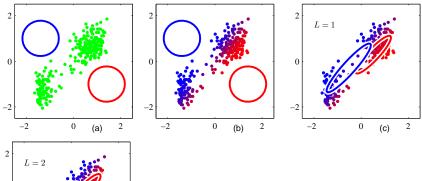
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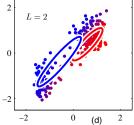


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Expectation-Maximization for GMMs

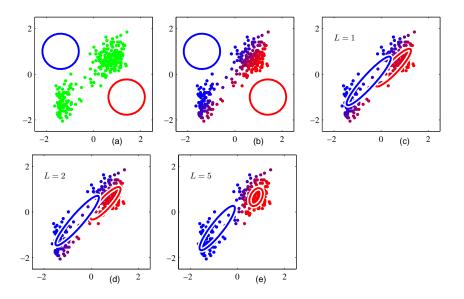




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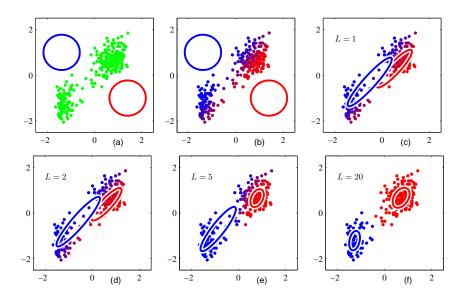
Expectation-Maximization for GMMs



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Expectation-Maximization for GMMs



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- Each step is more computationally intense than with K-Means too.
- So, one commonly computes K-Means first and then initializes EM from the resulting clusters.
- Care must be taken to avoid singularities in the MLE solution.
- There will generally be multiple local maxima of the likelihood function and EM is not guaranteed to find the largest of these.

Given a GMM, the goal is to maximize the likelihood function with respect to the parameters (the means, the covarianes, and the mixing coefficients).

() Initialize the means, μ_k , the covariances, Σ_k , and mixing coefficients, π_k . Evaluate the initial value of the log-likelihood.

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Given a GMM, the goal is to maximize the likelihood function with respect to the parameters (the means, the covarianes, and the mixing coefficients).

- Initialize the means, μ_k , the covariances, Σ_k , and mixing coefficients, π_k . Evaluate the initial value of the log-likelihood.
- **2** E-Step Evaluate the responsibilities using the current parameter values:

$$\gamma(z_k) = \frac{\pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

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M-Step Update the parameters using the current responsibilities

$$\boldsymbol{\mu}_{k}^{\text{new}} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{x}_{n}$$
⁽²⁹⁾

$$\boldsymbol{\Sigma}_{k}^{\mathsf{new}} = \frac{1}{N_{k}} \sum_{n=1}^{N} \gamma(z_{nk}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{\mathsf{new}}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{\mathsf{new}})^{\mathsf{T}}$$
(30)

$$\pi_k^{\text{new}} = \frac{N_k}{N} \tag{31}$$

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Evaluate the log-likelihood

$$\ln p(\mathbf{X}|\boldsymbol{\mu}^{\mathsf{new}}, \boldsymbol{\Sigma}^{\mathsf{new}}, \boldsymbol{\pi}^{\mathsf{new}}) = \sum_{n=1}^{N} \ln \left[\sum_{k=1}^{K} \pi_{k}^{\mathsf{new}} \mathcal{N}(\mathbf{x}_{n} | \boldsymbol{\mu}_{k}^{\mathsf{new}}, \boldsymbol{\Sigma}_{k}^{\mathsf{new}}) \right]$$
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Ocheck for convergence of either the parameters of the log-likelihood. If the convergence is not satisfied, set the parameters:

$$\boldsymbol{\mu} = \boldsymbol{\mu}^{\mathsf{new}} \tag{34}$$

$$\Sigma = \Sigma^{\text{new}} \tag{35}$$

$$\pi = \pi^{\text{new}}$$
 (36)

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and goto step 2.

• The goal of EM is to find maximum likelihood solutions for models having latent variables.

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- Denote the set of all model parameters as θ , and so the log-likelihood function is

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 - Even if the joint distribution $p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})$ belongs to the exponential family, the marginal $p(\mathbf{X}|\boldsymbol{\theta})$ typically does not.

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 - Even if the joint distribution $p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})$ belongs to the exponential family, the marginal $p(\mathbf{X}|\boldsymbol{\theta})$ typically does not.
- If, for each sample x_n we were given the value of the latent variable z_n, then we would have a complete data set, {X, Z}, with which maximizing this likelihood term would be straightforward.

• However, in practice, we are not given the latent variables values.

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- In the E-Step, we use the current parameter values θ^{old} to find the posterior distribution of the latent variables given by $p(\mathbf{Z}|\mathbf{X}, \theta^{\text{old}})$.
- This posterior is used to define the expectation of the complete-data log-likelihood, denoted Q(θ, θ^{old}), which is given by

$$Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{\mathsf{old}}) = \sum_{\mathbf{Z}} p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\theta}^{\mathsf{old}}) \ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\theta})$$
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• Then, in the M-step, we revise the parameters to θ^{new} by maximizing this function:

$$\boldsymbol{\theta}^{\mathsf{new}} = \arg \max_{\boldsymbol{\theta}} \mathcal{Q}(\boldsymbol{\theta}, \boldsymbol{\theta}^{\mathsf{old}})$$
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• Note that the log acts directly on the joint distribution $p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\theta})$ and so the M-step maximization will likely be tractable.