## CSE 632: Analysis of Algorithms II

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Lecture 24 (11/17/2017): Sketching and Streaming (II)

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## **24.1** Computing $F_2$

**Problem description:** Given data stream  $\langle i_1, \ldots, i_n \rangle$ , where  $i_t \in [m] \forall t \in [n]$ ; and let  $\mathcal{X} = (x_1, x_2, \ldots, x_m)$  s.t.  $x_i$  denote the number of times that i appears in the stream. The goal is to compute  $F_2$  — the  $\ell_2$ -norm of  $\mathcal{X}$ , i.e.,  $f = \|\mathcal{X}\|_2^2 = \sum_{i \in [m]} x_i^2$ , using only  $\operatorname{poly}(1/\epsilon, 1/\delta, \log n, \log m)$  storage.

**Analysis:** We choose a function h uniformly at random from  $\mathcal{H} = \{h : [m] \mapsto \{-1, +1\}\}$ , where  $\mathcal{H}$  is the set of all function mapping from [m] to  $\{-1, +1\}$ . And our estimator for f is as follows:

Algorithm 1 Estimator for f1:  $Z \leftarrow 0$ ; 2: for every  $i_t$  comes do 3:  $Z \leftarrow Z + h(i_t)$ ; 4: Output  $\hat{f} = Z^2$ ;

The first observation is:

$$Z = \sum_{i=1}^{m} x_i h(i)$$

And its expectation (w.r.t. the choice of h) is

$$\mathbb{E}[Z^2] = \mathbb{E}[\sum_{i=1}^m \sum_{j=1}^m x_i x_j h(i) h(j)] = \sum_{i=1}^m \sum_{j=1}^m x_i x_j \mathbb{E}[h(i) h(j)] = \sum_{i=1}^m x_i^2 = f(i) \mathbb{E}[Z^2]$$

i.e.,  $Z^2$  is an unbiased estimator for f. Therefore, if  $Z^2$  has a small variance, we can give an accurate estimate for  $F_2$ . By definition, the variance of  $Z^2$  is  $\operatorname{Var}[Z^2] = \mathbb{E}[Z^4] - \mathbb{E}^2[Z^2]$ , where

$$\mathbb{E}[Z^4] = \mathbb{E}\left[\sum_{\substack{i,j,k,l \in [m] \\ i,j,k,l \in [m]}} x_i x_j x_k x_l h(i) h(j) h(k) h(l)\right]$$
$$= \sum_{\substack{i,j,k,l \in [m] \\ i \in [m]}} x_i^4 + \binom{4}{2} \sum_{\substack{i,j \in [m] \\ i,j \in [m]}} x_i^2 x_j^2$$
(24.1)

The third equality is because  $\mathbb{E}[h(i)h(j)h(k)h(l)]$  will be zero except for the case  $\mathbb{E}[h(i)^4]$  or  $\mathbb{E}[h(i)^2h(j)^2]$ . Thus,

$$\begin{aligned} \operatorname{Var}[Z^2] &= \sum_{i \in [m]} x_i^4 + 6 \sum_{i,j \in [m]} x_i^2 x_j^2 - (\sum_{i \in [m]} x_i^2)^2 \\ &= 4 \sum_{i,j \in [m]} x_i^2 x_j^2 \\ &\leq 2(\sum_{i \in [m]} x_i^2)^2 = 2f^2 \end{aligned}$$

To clear notations, let  $Y = Z^2$ , then  $\mathbb{E}[Y] = f$ ,  $\operatorname{Var}[Y] \leq 2f^2$ . Now similar to the last lecture, we call each run of Algorithm 1 as a "single experiment", and run  $k = \lceil \frac{2}{\delta \epsilon^2} \rceil$  experiments independently. Denote the output of *i*-th run as  $Y_i$ , and define  $Y_0 = \frac{1}{k}(Y_1 + Y_2 + \cdots + Y_k)$ , then we have

$$\mathbb{E}[Y_0] = f, \quad \operatorname{Var}[Y_0] \le \frac{2f^2}{k}$$

By Chebyshev's inequality,

$$\Pr[|Y_0 - \mathbb{E}[Y_0]| \ge \epsilon f] \le \frac{\operatorname{Var}[Y_0]}{\epsilon^2 f^2} = \frac{2}{k\epsilon^2} \le \delta$$

So we get a  $(\epsilon, \delta)$ -approximation with  $O\left(\frac{1}{\delta\epsilon^2}\right)$  memory cost. Actually the  $O(1/\delta)$  factor can be further reduced by then so-called "median trick": replace each single experiment with the mean output of  $O(1/\epsilon^2)$  experiments, and make  $k = O(\log(1/\delta))$  such means (i.e., run  $O\left(\frac{1}{\epsilon^2}\log\frac{1}{\delta}\right)$  experiments in total). Then we output the median of these  $O(1/\epsilon^2)$  means as the final result. Specifically, let  $\bar{Y}^t$  be the mean of the *t*-th  $6/\epsilon^2$  experiment repetitions, and consider the random variables  $W^t = \mathbb{I}\{|\bar{Y}^t - f| > \epsilon f\}(t = 1, \ldots, k)$ : by Chebyshev's inequality, each  $W^t$  is a Bernoulli random variable with  $\Pr[W^t = 1] \leq 1/3$ . So if the median of all  $\bar{Y}^t$ s is at least  $\epsilon f$  far away from f, then at least half of all  $W^t$ s are 1, of which the probability, by a standard Chernoff bound, decreases exponentially in k.

There's only one problem left: how do we sample the random function h? The function set  $\mathcal{H}$  is of size  $2^m$ , sampling uniformly random from  $\mathcal{H}$  will require O(m) random bits, which is unacceptable. But from the derivation of variance (24.1), we can see that h only needs to be 4-wise independent:

**Definition 24.1 (4-wise independent function)** Let  $\mathcal{H}$  be a family of functions from A to B,  $\mathcal{H}$  is 4-wise independent if  $\forall$  4 distinct elements  $a, b, c, d \in A$ , and 4 values  $v_a, v_b, v_c, v_d \in B$ , we have

$$\Pr_{h \in \mathcal{H}}[h(a) = v_a, h(b) = v_b, h(c) = v_c, h(d) = v_d] =$$
$$\Pr_{h \in \mathcal{H}}[h(a) = v_a] \Pr_{h \in \mathcal{H}}[h(b) = v_b] \Pr_{h \in \mathcal{H}}[h(c) = v_c] \Pr_{h \in \mathcal{H}}[h(d) = v_d]$$

**Example 24.2 (4-wise independent function)** Let q be a prime, and define  $g_{a,b,c,d} : [m] \mapsto \mathbb{F}_{2^q}$  as follows:

$$g_{a,b,c,d}(x) = ax^3 + bx^2 + cx + d$$

here  $a, b, c, d \in \mathbb{F}_{2^q}$ . Now let  $h_{a,b,c,d}(x) = 2(g_{a,b,c,d}(x) \mod 2) - 1$ , then  $\mathcal{H} = \{h_{a,b,c,d} : a, b, c, d \in \mathbb{F}_{2^q}\}$  is a 4-wise independ function family.

## 24.2 Summary

From these two lectures we conclude that  $F_0$  and  $F_2$  can be estimated with  $poly(1/\epsilon, log(1/\delta), log n, log m)$ memory space. And the method we give can even work with stream in more general forms:  $\langle (i_1, \Delta_1), (i_2, \Delta_2), \ldots, (i_t, \Delta_t) \rangle$ , such that  $x_i = \sum_{t:i_t=i} \Delta_t$ . Actually for any  $F_p$  with  $p \in [0, 2]$ , we can obtain a  $(\epsilon, \delta)$ -approximation with  $poly(1/\epsilon, log(1/\delta), log n, log m)$  space. But for  $F_p$  where p > 2, we always need  $n^{\Omega(1)}$  space.